Extrema of Branching Brownian Motion in \mathbb{R}^d

Yujin Kim (Courant Institute, NYU)

MSRI Program Associate Short Talks, September 2021 Joint works with Julien Berestycki, Bastien Mallein, Eyal Lubetzky, and Ofer Zeitouni. Paradigmatic model in the study of extrema of log-correlated fields– see Ofer's talk from 9/1.

- Random matrix theory
- Statistics of Riemann zeta function
- GFF and GMC
- Cover times of random walks
- PDEs (F-KPP equation)
- ...

Model Definition

Fix the dimension $d \ge 1$.

• Start: a single particle v at 0 performs Brownian motion in \mathbb{R}^d (iid 1d BM's in each coordinate)

 $X_{s}^{(v)} = (W_{s}^{1,(v)}, \dots, W_{s}^{d,(v)}) \in \mathbb{R}^{d}$

Model Definition

Fix the dimension $d \ge 1$.

• Start: a single particle v at 0 performs Brownian motion in \mathbb{R}^d (iid 1d BM's in each coordinate)

$$X_s^{(v)} = (W_s^{1,(v)}, \ldots, W_s^{d,(v)}) \in \mathbb{R}^d$$

• After exp(1) distributed time, the particle splits into two particles, which carry on independently.

Model Definition

Fix the dimension $d \ge 1$.

• Start: a single particle v at 0 performs Brownian motion in \mathbb{R}^d (iid 1d BM's in each coordinate)

$$X_s^{(v)} = (W_s^{1,(v)}, \ldots, W_s^{d,(v)}) \in \mathbb{R}^d$$

- After $\exp(1)$ distributed time, the particle splits into two particles, which carry on independently.
- Repeat.

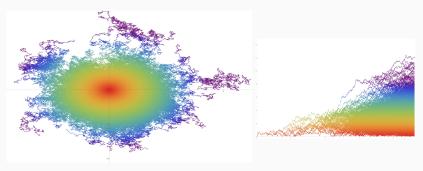


Fig: 2D BBM (left) and its modulus as a function of time (right)

Fix the dimension $d \ge 1$.

• Start: a single particle v at 0 performs Brownian motion in \mathbb{R}^d (iid 1d BM's in each coordinate)

$$X^{(v)}_s = (W^{1,(v)}_s, \ldots, W^{d,(v)}_s) \in \mathbb{R}^d$$

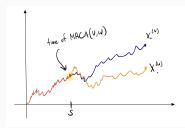
- After an $\exp(1)$ distributed time, the particle splits into two particles, which perform independent (*d*-dim) BM's from that point forward and carry their own exponential clocks
- Repeat.

Notation.

- Let $N_t :=$ set of particles at time t.
- $X_{\cdot}^{(v)} :=$ the *d*-dim BM path of *v*.

Quick facts.

- $\mathbb{E}[N_t] = e^t$
- For coordinate indices $i, j \in \{1, ..., d\}$, $\operatorname{Cov}(X_t^{(v)}(i), X_t^{(w)}(j)) = \text{time of the most}$ recent common ancestor of v and w



We will write $R_t^* := \max_{v \in N_t} \|X_t^{(v)}\|$.

We will write $R_t^* := \max_{v \in N_t} \|X_t^{(v)}\|$.

Dimension 1 results:

• Bramson: found centering term $m_t(1) = \sqrt{2}t - \frac{3}{2\sqrt{2}}\log t$, and showed $R_t^* - m_t(1)$ converges in law via connection to F-KPP equation.

Theorem (Bramson '77, '83)

 $\lim_{t\to\infty} \mathbb{P}(R_t^* - m_t(1) \le y) = w(y)$, where $\frac{1}{2}w'' + \sqrt{2}w' + w^2 - w = 0$.

We will write $R_t^* := \max_{v \in N_t} \|X_t^{(v)}\|$.

Dimension 1 results:

• Bramson: found centering term $m_t(1) = \sqrt{2}t - \frac{3}{2\sqrt{2}}\log t$, and showed $R_t^* - m_t(1)$ converges in law via connection to F-KPP equation.

Theorem (Bramson '77, '83)

 $\lim_{t\to\infty} \mathbb{P}(R_t^* - m_t(1) \le y) = w(y)$, where $\frac{1}{2}w'' + \sqrt{2}w' + w^2 - w = 0$.

• Idea from physics: the distribution of the max. of log-correlated fields should still resemble classical (iid) extreme value distributions

We will write $R_t^* := \max_{v \in N_t} \|X_t^{(v)}\|$.

Dimension 1 results:

• Bramson: found centering term $m_t(1) = \sqrt{2}t - \frac{3}{2\sqrt{2}}\log t$, and showed $R_t^* - m_t(1)$ converges in law via connection to F-KPP equation.

Theorem (Bramson '77, '83)

 $\lim_{t\to\infty} \mathbb{P}(R_t^* - m_t(1) \le y) = w(y)$, where $\frac{1}{2}w'' + \sqrt{2}w' + w^2 - w = 0$.

- Idea from physics: the distribution of the max. of log-correlated fields should still resemble classical (iid) extreme value distributions
- Lalley-Selke: characterization of the limiting law as a *randomly shifted Gumbel*, where the random shift comes from the behavior of the particles at the beginning of the process (more on this later).

Theorem (Lalley-Selke '87)

There exists a random variable Z, called the derivative martingale, and a constant C>0 such that

$$\lim_{t\to\infty}\mathbb{P}(R^*_t-m_t(1)\leq y)=\mathbb{E}\Big[\exp\big(-e^{-\sqrt{2}(y-CZ)}\big)\Big]\,,\,\,\forall y\in\mathbb{R}\,.$$

What about the point process formed by all particles "near the maximum": the extremal point process?

$$\mathcal{E}_t := \sum_{v \in N_t} \delta_{\|X_t^{(v)}\| - m_t(1)}.$$

- The limit was identified independently by Aidekon-Berestycki-Brunet-Shi and Arguin-Bovier-Kistler as a randomly shifted, decorated Poisson point process
- The shift is given by the log of the derivative martingale Z.

Theorem (ABBS/ABK 2011)

$$\mathcal{E} := \sum_{i=1}^{\infty} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{\eta_i + \xi_j + \log Z} \,.$$



What about the point process formed by all particles "near the maximum": the extremal point process?

$$\mathcal{E}_t := \sum_{v \in N_t} \delta_{\|X_t^{(v)}\| - m_t(1)}$$

- The limit was identified independently by Aidekon-Berestycki-Brunet-Shi and Arguin-Bovier-Kistler as a randomly shifted, decorated Poisson point process
- The shift is given by the log of the derivative martingale Z.

Theorem (ABBS/ABK 2011)

$$\mathcal{E} := \sum_{i=1}^{\infty} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{\eta_i + \xi_j + \log Z} \,.$$



What about the point process formed by all particles "near the maximum": the extremal point process?

$$\mathcal{E}_t := \sum_{v \in N_t} \delta_{\|X_t^{(v)}\| - m_t(1)}$$

- The limit was identified independently by Aidekon-Berestycki-Brunet-Shi and Arguin-Bovier-Kistler as a randomly shifted, decorated Poisson point process
- The shift is given by the log of the derivative martingale Z.

Theorem (ABBS/ABK 2011)

$$\mathcal{E} := \sum_{i=1}^{\infty} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{\eta_i + \xi_j + \log Z} \,.$$



What about the point process formed by all particles "near the maximum": the extremal point process?

$$\mathcal{E}_t := \sum_{v \in N_t} \delta_{\|X_t^{(v)}\| - m_t(1)}.$$

- The limit was identified independently by Aidekon-Berestycki-Brunet-Shi and Arguin-Bovier-Kistler as a randomly shifted, decorated Poisson point process
- The shift is given by the log of the derivative martingale Z.

Theorem (ABBS/ABK 2011)

$$\mathcal{E} := \sum_{i=1}^{\infty} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{\eta_i + \xi_j + \log Z} \,.$$



What about the point process formed by all particles "near the maximum": the extremal point process?

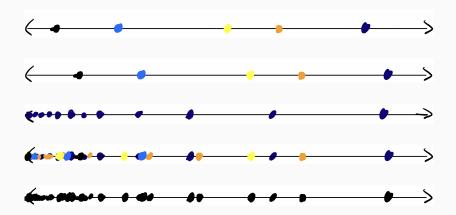
$$\mathcal{E}_t := \sum_{v \in N_t} \delta_{\|X_t^{(v)}\| - m_t(1)}.$$

- The limit was identified independently by Aidekon-Berestycki-Brunet-Shi and Arguin-Bovier-Kistler as a randomly shifted, decorated Poisson point process
- The shift is given by the log of the derivative martingale Z.

Theorem (ABBS/ABK 2011)

$$\mathcal{E} := \sum_{i=1}^{\infty} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{\eta_i + \xi_j + \log Z} \,.$$





• Lots of further work: extending convergence to include genealogical info (Bovier-Hartung 2016), structure of extreme level sets (Cortines-Hartung-Louidor 2017), etc.

Does the maximum modulus R_t^* in $d \ge 2$ also converge in distribution, after re-centering, to some perturbation of a classical extreme value distribution?

Does the maximum modulus R_t^* in $d \ge 2$ also converge in distribution, after re-centering, to some perturbation of a classical extreme value distribution? Yes!

Theorem (K.-Lubetzky-Zeitouni 2021)

Let $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$. There exists a non-degenerate, positive random variable Z_∞ and a constant $C_d > 0$ such that

$$\lim_{t\to\infty} \mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\exp\Big(-C_d e^{-\sqrt{2}(y-\frac{1}{\sqrt{2}}\log Z_\infty)}\Big)\Big], \ \forall y \in \mathbb{R}$$

Does the maximum modulus R_t^* in $d \ge 2$ also converge in distribution, after re-centering, to some perturbation of a classical extreme value distribution? Yes!

Theorem (K.-Lubetzky-Zeitouni 2021)

Let $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$. There exists a non-degenerate, positive random variable Z_∞ and a constant $C_d > 0$ such that

$$\lim_{t\to\infty} \mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\exp\Big(-C_d e^{-\sqrt{2}(y-\frac{1}{\sqrt{2}}\log Z_\infty)}\Big)\Big], \ \forall y \in \mathbb{R}$$

Time permitting, we will also discuss forthcoming work describing the limiting law of the extremal point process as a randomly shifted, decorated PPP on \mathbb{R}^d .

Does the maximum modulus R_t^* in $d \ge 2$ also converge in distribution, after re-centering, to some perturbation of a classical extreme value distribution? Yes!

Theorem (K.-Lubetzky-Zeitouni 2021)

Let $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$. There exists a non-degenerate, positive random variable Z_∞ and a constant $C_d > 0$ such that

$$\lim_{t\to\infty} \mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\exp\Big(-C_d e^{-\sqrt{2}(y-\frac{1}{\sqrt{2}}\log Z_\infty)}\Big)\Big], \ \forall y \in \mathbb{R}$$

Time permitting, we will also discuss forthcoming work describing the limiting law of the extremal point process as a randomly shifted, decorated PPP on \mathbb{R}^d .

Previous work:

- Biggins '95: $R_t^*/\sqrt{2}t \rightarrow 1$ a.s.
- Mallein 2015: $\{R_t^* m_t\}_{t>0}$ is tight.

• Recall that the norm of a *d*-dimensional Brownian motion is a *d*-dimensional Bessel process.

• Recall that the norm of a *d*-dimensional Brownian motion is a *d*-dimensional Bessel process.

SDE for Bessel process

$$\mathrm{d}R_t = \frac{d-1}{2R_t}\mathrm{d}t + \mathrm{d}W_t$$

• Note: Markov process on \mathbb{R} , but not shift-invariant.

• Recall that the norm of a *d*-dimensional Brownian motion is a *d*-dimensional Bessel process.

SDE for Bessel process

$$\mathrm{d}R_t = \frac{d-1}{2R_t}\mathrm{d}t + \mathrm{d}W_t$$

- Note: Markov process on \mathbb{R} , but not shift-invariant.
- We study the *d*-dim. branching Bessel process $\{R_s^{(v)}\}_{s>0,v\in N_s}$.

• Recall that the norm of a *d*-dimensional Brownian motion is a *d*-dimensional Bessel process.

SDE for Bessel process

$$\mathrm{d}R_t = rac{d-1}{2R_t}\mathrm{d}t + \mathrm{d}W_t$$

- Note: Markov process on \mathbb{R} , but not shift-invariant.
- We study the *d*-dim. branching Bessel process $\{R_s^{(v)}\}_{s>0,v\in N_s}$.
- The Girsanov transform gives us the R-N derivative of the law P^R of a Bessel process w.r.t. the law P^W of Brownian motion W.

• Recall that the norm of a *d*-dimensional Brownian motion is a *d*-dimensional Bessel process.

SDE for Bessel process

$$\mathrm{d}R_t = \frac{d-1}{2R_t}\mathrm{d}t + \mathrm{d}W_t$$

- Note: Markov process on \mathbb{R} , but not shift-invariant.
- We study the *d*-dim. branching Bessel process $\{R_s^{(v)}\}_{s>0,v\in N_s}$.
- The Girsanov transform gives us the R-N derivative of the law P^R of a Bessel process w.r.t. the law P^W of Brownian motion W.

Girsanov transform from Bessel to BM

$$\mathrm{d}P^{R}\big|_{\mathcal{F}_{t}} = \underbrace{\left(\frac{W_{t}}{W_{0}}\right)^{\frac{d-1}{2}}}_{\substack{\text{start/endpoint} \\ \text{dependence}}} \underbrace{\exp\left(\int_{0}^{t} \frac{c_{d}}{W_{u}^{2}} \mathrm{d}u\right) \mathbb{1}_{\{W_{u} > 0, \ u \in [0, t]\}}}_{pathwise \ dependence} \mathrm{d}P^{W}\big|_{\mathcal{F}_{t}},$$

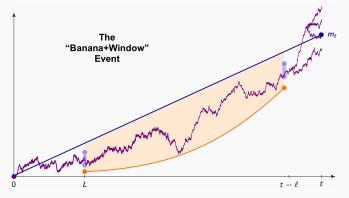
where $c_d > 0$ for $d \ge 3$ and $c_d < 0$ for d = 2.

Trajectories of the extremal particles

• Let L, ℓ be parameters that we send to infinity after t (think: constants wrt t).

Trajectories of the extremal particles

• Let L, ℓ be parameters that we send to infinity after t (think: constants wrt t).



Extremal particle trajectories

W.h.p., all particles $v \in N_t$ that reach height m_t at time t did the following:

- at time L, passed through the "window" $\sqrt{2}L [L^{1/6}, L^{2/3}]$
- at time $t-\ell$, passed through the "window" $\frac{m_t}{t}(t-\ell)-[\ell^{1/3},\ell^{2/3}]$
- on $[L, t \ell]$, stayed within the "banana"

Say we have a particle $\nu \in \mathit{N}_L$ such that $\mathit{R}_L^{(\nu)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

Say we have a particle $v\in \mathit{N}_L$ such that $\mathit{R}_L^{(v)}=r\in\sqrt{2}L-[\mathit{L}^{1/6},\mathit{L}^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

 $(Markov) = \mathbb{P}_r(R_{t-L}^* > m_t, Banana+Window event)$

$$(\mathsf{Markov}) = \mathbb{P}_r(R_{t-L}^* > m_t, \mathsf{Banana+Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$

$$\begin{aligned} \mathsf{(Markov)} &= \mathbb{P}_r(R_{t-L}^* > m_t, \; \mathsf{Banana} + \mathsf{Window \; event}) \\ &= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big) \\ &\sim e^{t-L-\ell} \mathbb{P}_r(BW(u)) \\ &\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} \,. \end{aligned}$$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$\begin{aligned} \mathsf{(Markov)} &= \mathbb{P}_r(R_{t-L}^* > m_t, \; \mathsf{Banana} + \mathsf{Window \; event}) \\ &= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big) \\ &\sim e^{t-L-\ell} \mathbb{P}_r(BW(u)) \\ &\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} \,. \end{aligned}$$

Remarks:

(1) The fact that this asymptotic has **no** *t*-dependence means that $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$ is the right centering term!

Say we have a particle $v\in \mathit{N}_L$ such that $\mathit{R}_L^{(v)}=r\in\sqrt{2}L-[\mathit{L}^{1/6},\mathit{L}^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$\begin{aligned} \mathsf{Markov} &= \mathbb{P}_r(R^*_{t-L} > m_t, \; \mathsf{Banana+Window\; event}) \\ &= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big) \\ &\stackrel{\mathbf{v}}{\sim} e^{t-L-\ell} \mathbb{P}_r(BW(u)) \\ &\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L-r) e^{-\sqrt{2}(\sqrt{2}L-r)} \,. \end{aligned}$$

Remarks:

- (1) The fact that this asymptotic has no *t*-dependence means that $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$ is the right centering term.
- (2) A lot actually goes into showing the highlighted \sim :

Say we have a particle $v \in N_L$ such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$\begin{aligned} \mathsf{Markov} &= \mathbb{P}_r(R^*_{t-L} > m_t, \; \mathsf{Banana} + \mathsf{Window \; event}) \\ &= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big) \\ &\stackrel{\mathbf{\sim}}{\sim} e^{t-L-\ell} \mathbb{P}_r(BW(u)) \\ &\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)}. \end{aligned}$$

Remarks:

- (1) The fact that this asymptotic has no *t*-dependence means that $m_t := \sqrt{2}t + \frac{d-4}{2\sqrt{2}}\log t$ is the right centering term.
- (2) A *lot* actually goes into showing the highlighted ~:
 We show that, on the Banana+Window event, at most one u ∈ N_{t-L-ℓ} produces a descendant that exceeds height m_t. Done via a "modified" second moment method, inspired by the work of Bramson-Ding-Zeitouni (2013) establishing convergence of the re-centered maximum of the 2d discrete GFF.

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

$$(\mathsf{Markov}) = \mathbb{P}_r(R^*_{t-L} > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} =: C_d f(r).$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R^*_{t-L} > m_t, \mathsf{Banana+Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell}\mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}}(\sqrt{2}L-r)e^{-\sqrt{2}(\sqrt{2}L-r)} =: C_d f(r).$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R_{t-L}^* > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} =: C_d f(r).$$

$$\mathbb{P}(R_t^* \leq m_t + y) = \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \leq m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \big(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\big)\Big]$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R_{t-L}^* > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} =: C_d f(r).$$

$$\begin{split} \mathbb{P}(R_t^* \le m_t + y) &= \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \le m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\right)\Big] \\ &\approx \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - C_d f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big] \end{split}$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R^*_{t-L} > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L-r)} =: C_d f(r).$$

$$\mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \le m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\right)\Big]$$
$$\approx \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - C_d f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big] \approx \mathbb{E}\Big[\exp\Big(-C_d \sum_{v \in N_L} f(R_L^{(v)})e^{-y\sqrt{2}}\Big)\Big]$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R^*_{t-L} > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L-r)} =: C_d f(r).$$

$$\mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \le m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\right)\Big]$$
$$\approx \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - C_d f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big] \approx \mathbb{E}\Big[\exp\Big(-C_d \sum_{v \in N_L} f(R_L^{(v)})e^{-y\sqrt{2}}\Big)\Big]$$

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R_{t-L}^* > m_t, \mathsf{Banana+Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} =: C_d f(r).$$

Proof of main result:

$$\mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \le m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\right)\Big]$$
$$\approx \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - C_d f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big] \approx \mathbb{E}\Big[\exp\Big(-C_d \sum_{v \in N_L} f(R_L^{(v)})e^{-y\sqrt{2}}\Big)\Big]$$
$$\rightarrow \mathbb{E}\Big[\exp\Big(-C_d Z_\infty e^{-y\sqrt{2}}\Big)\Big]$$

 Note the LHS has no L dependence, while the RHS has no t dependence ⇒ both sides converge!

Say we have a particle
$$v \in N_L$$
 such that $R_L^{(v)} = r \in \sqrt{2}L - [L^{1/6}, L^{2/3}].$

What is the probability that v produces a descendant in N_t that exceeds height m_t ?

$$(\mathsf{Markov}) = \mathbb{P}_r(R_{t-L}^* > m_t, \mathsf{Banana} + \mathsf{Window event})$$
$$= \mathbb{P}_r\Big(\bigcup_{u \in N_{t-L-\ell}} BW(u)\Big)$$
$$\sim e^{t-L-\ell} \mathbb{P}_r(BW(u))$$
$$\sim C_d r^{-\frac{d-1}{2}} (\sqrt{2}L - r) e^{-\sqrt{2}(\sqrt{2}L - r)} =: C_d f(r).$$

$$\mathbb{P}(R_t^* \le m_t + y) = \mathbb{E}\Big[\prod_{v \in N_L} \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* \le m_t + y)\Big] = \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - \mathbb{P}_{R_L^{(v)}}(R_{t-L}^* > m_t + y)\right)\Big]$$
$$\approx \mathbb{E}\Big[\prod_{v \in N_L} \left(1 - C_d f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big] \approx \mathbb{E}\Big[\exp\left(-C_d \sum_{v \in N_L} f(R_L^{(v)})e^{-y\sqrt{2}}\right)\Big]$$
$$\rightarrow \mathbb{E}\Big[\exp\left(-C_d Z_\infty e^{-y\sqrt{2}}\right)\Big]$$

- Note the LHS has no L dependence, while the RHS has no t dependence ⇒ both sides converge!
- Z_{∞} is the limit in distribution of $\sum_{v \in N_L} f(R_L^{(v)})$. Convergence in \mathbb{P} forthcoming work with J. Berestycki, B. Mallein.

Interpretation of Z_{∞} by Lalley-Selke

Theorem (Berestycki-K.-Lubetzky-Mallein-Zeitouni 2021)

$$\mathbb{P}(R_t^* - m_t \leq y \mid \mathcal{F}_L) \xrightarrow[L \to \infty]{t \to \infty} \exp\left(-C_d Z_\infty e^{-\sqrt{2}y}\right)$$

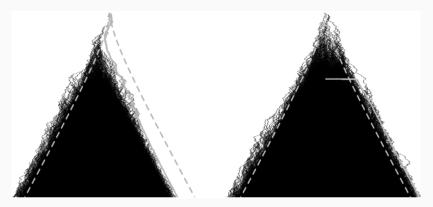


Figure 2: Left: initial particles veer to the left. Right: initial particles veer far to the right. In both pictures, we see how the initial behavior permanently shifts the maximum. (Image from Éric Brunet).

Main Results: The Extremal Point Process in \mathbb{R}^d

- Stasiński, Berestycki, and Mallein (2020) constructed a random measure $D_{\infty}(\cdot)$ on \mathbb{S}^{d-1} , which is the a.s., a.e. limit of an "angular" derivative martingale. So, Z_{∞} and $D_{\infty}(\mathbb{S}^{d-1})$ have the same law.
- D_{∞} is a.s. absolutely continuous wrt Leb.
- It turns out that ∑_{v∈NL} f(R^(v)_L) converges in probability to c_dD_∞(S^{d-1}), for some constant c_d > 0.
- Let the point process $\sum_{i=1}^{\infty} \delta_{(\eta_i,\theta_i)}$ on $\mathbb{R} \times \mathbb{S}^{d-1}$ have law $\mathsf{PPP}\Big(\overline{C}_d D_{\infty}(\mathbb{S}^{d-1}) e^{-\sqrt{2}x} \mathrm{d}x \times \frac{D_{\infty}(\theta)}{D_{\infty}(\mathbb{S}^{d-1})} \mathrm{d}\theta\Big).$
- Let $\mathcal{E}_t := \sum_{v \in N_t} \delta(\|X_t\| m_t, X_t/\|X_t\|)$ denote the extremal point process of BBM in dimension d.

Theorem

Let $\{\mathcal{D}^{(i)}\}_{i \in \mathbb{N}}$ be a collection of iid point processes with the same law as the decorations from the 1D BBM case. Then \mathcal{E}_t converges in law as $t \to \infty$ to

$$\mathcal{E} := \sum_{i \in \mathbb{N}} \sum_{\xi_j \in \mathcal{D}^{(i)}} \delta_{(\eta_i + \xi_j, \theta_i)}$$



